

Probability of fixation and evolution of cooperation

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Thanks

Collaborators

Véronique Ladret

Philippe Lahaie

Samuel Langevin

David Lasalle-Ialongo

John Wakeley

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Philippe Lahaie
Samuel Langevin
David Lasalle-Ialongo
John Wakeley

Competitors

Martin Nowak
François Rousset
and many others



1. Examples of cooperation











Cooperation is widespread...

but emergence of cooperation is problematic!

- ▶ A game-theoretic framework based on pairwise interactions is a first step toward a better understanding of a complex phenomenon.
- ▶ This framework can provide clues about conditions that would favor the evolution of cooperation.

2. Prisoner's Dilemma (PD)



Payoff Matrix

<i>Cooperate</i>	<i>Reward</i>	Sucker's payoff
<i>Defect</i>	<i>Temptation</i>	<i>Punishment</i>
against	<i>Cooperate</i>	<i>Defect</i>

$$T > R > P > S$$

<i>C</i>	5	1
<i>D</i>	14	3
against	<i>C</i>	<i>D</i>

Iterated Prisoner's Dilemma (IPD)

- ▶ PD repeated n times between the same players with additive payoffs

Tit-for-Tat (A)	$a = Rn$	$b = S + P(n - 1)$
Always-Defect (B)	$c = T + P(n - 1)$	$d = Pn$
against	A	B

$$a > c > d > b \text{ for } n > \frac{T-P}{R-P}$$

<i>A</i>	50	28
<i>B</i>	41	30
against	<i>A</i>	<i>B</i>

for $n = 10$ in the previous example

Expected payoffs in an infinite population

- ▶ random pairwise interactions
- ▶ x : frequency of A

$$w_A(x) = ax + b(1 - x)$$

$$w_B(x) = cx + d(1 - x)$$

$$w_A(x) > w_B(x) \text{ if and only if } x > \frac{d-b}{a-b-c+d} = x^* \downarrow 0 \text{ as } n \uparrow \infty$$

3. Evolutionary dynamics in an infinite population

- ▶ reproduction with discrete, non-overlapping generations
- ▶ $s \geq 0$: intensity of selection with expected payoff as coefficient
- ▶ $x(t)$: frequency of A in offspring in generation t before selection

$$x(t+1) = \frac{x(t)(1 + sw_A(x(t)))}{1 + s\bar{w}(x(t))}$$

$$x(t+1) - x(t) = \frac{s(a - b - c + d)x(t)(1 - x(t))(x(t) - x^*)}{1 + s\bar{w}(x(t))}$$

$$x(t) \uparrow 1 \text{ if } x(0) > x^* \text{ and } x(t) \downarrow 0 \text{ if } x(0) < x^*$$

4. Probability of fixation in a finite population

- ▶ N parents chosen at random to produce the next generation (assumed large)
- ▶ (v_1, \dots, v_N) : proportions of offspring produced in large numbers (assumed exchangeable)
- ▶ $X(t)$: frequency of A in offspring in generation t before selection
- ▶ $X(0) = N^{-1}$

$$X(T) \rightarrow X(\infty) = X(0) + \sum_{t \geq 0} (X(t+1) - X(t))$$

$X(\infty) = 1$ with probability $u(s)$, and 0 otherwise

► $u(s) = E_s[X(\infty)]$: probability of ultimate fixation of A

► $u(0) = X(0)$

$$\begin{aligned}u(s) &= X(0) + \sum_{t \geq 0} E_s[X(t+1) - X(t)] \\&= u(0) + s(a - b - c + d) \sum_{t \geq 0} E_s \left[\frac{X(t)(1 - X(t))(X(t) - x^*)}{1 + s\bar{w}(X(t))} \right] \\&\approx u(0) + s(a - b - c + d) \sum_{t \geq 0} E_0[X(t)(1 - X(t))(X(t) - x^*)]\end{aligned}$$

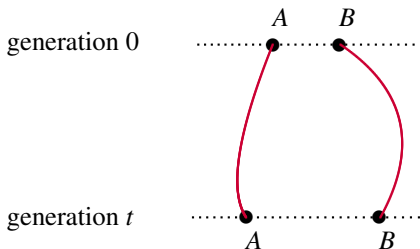
$u(s) > u(0)$ for $s > 0$ small : weak selection favors A replacing B

$$x^* < \frac{\sum_{t \geq 0} E_0[X(t)^2(1 - X(t))]}{\sum_{t \geq 0} E_0[X(t)(1 - X(t))]} = \tilde{x}$$

5. Generalized one-third law

$$\begin{aligned}\sum_{t \geq 0} E_0[X(t)(1 - X(t))] &= \sum_{t \geq 0} P_0(A, B \text{ in generation } t) \\ &= \sum_{t \geq 0} p_{22}(t) P_0(A, B \text{ in generation } 0) \\ &= \frac{X(0)(1 - X(0))}{1 - p_{22}}\end{aligned}$$

with $p_{22}(t) = p_{22}^t$ the probability that two offspring chosen at random in generation t descend from **two distinct ancestors** in generation 0

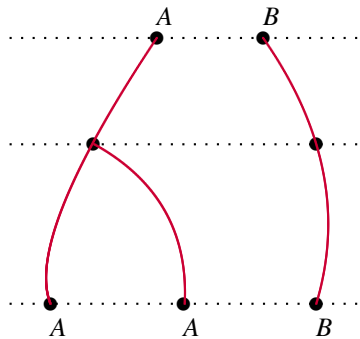
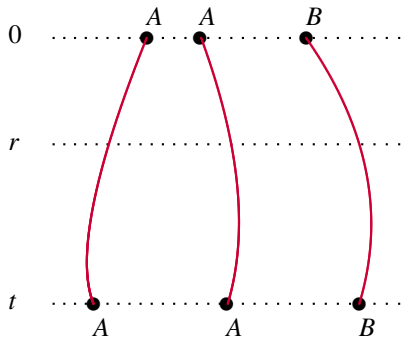


$$\begin{aligned}
\sum_{t \geq 0} E_0[X(t)^2(1 - X(t))] &= \sum_{t \geq 0} P_0(A, A, B \text{ in generation } t) \\
&= \sum_{t \geq 0} p_{33}(t) P_0(A, A, B \text{ in generation } 0) \\
&\quad + \sum_{t \geq 0} \frac{p_{32}(t)}{3} P_0(A, B \text{ in generation } 0)
\end{aligned}$$

with $p_{33}(t) = p_{33}^t$ the probability that three offspring chosen at random in generation t descend from **three distinct ancestors** in generation 0, and

$$p_{32}(t) = \sum_{r=0}^{t-1} p_{33}^{t-r-1} p_{32} p_{22}^r = p_{32} \left(\frac{p_{33}^t - p_{22}^t}{p_{33} - p_{22}} \right)$$

the probability that they descend from **two distinct ancestors** in generation 0.



$$\begin{aligned} \sum_{t \geq 0} E_0[X(t)^2(1 - X(t))] &= \frac{X(0)^2(1 - X(0))}{1 - p_{33}} + \frac{p_{32}X(0)(1 - X(0))}{3(1 - p_{22})(1 - p_{33})} \\ &\approx \frac{p_{32}X(0)(1 - X(0))}{3(1 - p_{22})(1 - p_{33})} \end{aligned}$$

and therefore

$$\tilde{x} \approx \frac{p_{32}}{3(1 - p_{33})} \leq \frac{1}{3}$$

with equality if and only if at most 2 lineages out of 3 coalesce at a time backward in time with probability 1, which characterizes the **Kingman coalescent** for a wide range of reproduction schemes as $N \rightarrow \infty$ with N generations as unit of time (Möhle 2000).

Λ -coalescent (e.g., Eldon and Wakeley 2006)

- ▶ $1 - N^{-\alpha}$: probability that every parent produces the same proportion N^{-1} of the offspring
- ▶ $N^{-\alpha}$: probability that a parent chosen at random produces a proportion ψ of offspring and every other parent a proportion $(1 - \psi)(N - 1)^{-1}$

$$\tilde{x} \approx \frac{1-\psi}{3-2\psi} < \frac{1}{3} \text{ if } \alpha < 1$$

which means a **more stringent condition** for cooperation to be favored in the case of a **highly skewed distribution** for the contribution of parents in offspring.

6. Projected average allelic effect

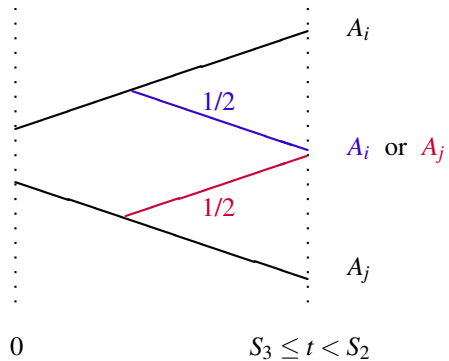
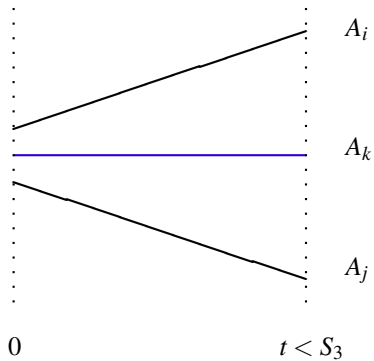
For any number of types A_1, \dots, A_n and any coefficient of selection w_{ij} for A_i in interaction with A_j , the first order effect of selection on the probability of fixation of A_i is

$$u'_i(0) = X_i(0)a_i(0)$$

where $a_i(0)$ is a projected average allelic effect on the selection coefficient given by

$$\begin{aligned} a_i(0) &= E_0(S_3) \left[\sum_k w_{ik} X_k(0) - \sum_j \sum_k w_{jk} X_j(0) X_k(0) \right] \\ &+ \left(\frac{E_0(S_2) - E_0(S_3)}{2} \right) \left[w_{ii} - \sum_j w_{ji} X_j(0) \right] \\ &+ \left(\frac{E_0(S_2) - E_0(S_3)}{2} \right) \left[\sum_j w_{ij} X_j(0) - \sum_j w_{jj} X_j(0) \right] \end{aligned}$$

with $E_0(S_j)$ being the expected number of generations with j lineages.



Symmetric case $w_{ij} = w_{ji}$

$$a_i(0) = E_0(S_3) \left[\sum_k w_{ik} X_k(0) - \sum_j \sum_k w_{jk} X_j(0) X_k(0) \right] \\ + \left(\frac{E_0(S_2) - E_0(S_3)}{2} \right) \left[w_{ii} - \sum_j w_{jj} X_j(0) \right]$$

with **autozygous pairs having the same weight as allozygous pairs**, leading to the one-third law of evolution, in the domain of application of the Kingman coalescent, and **a lighter weight than allozygous pairs** otherwise.

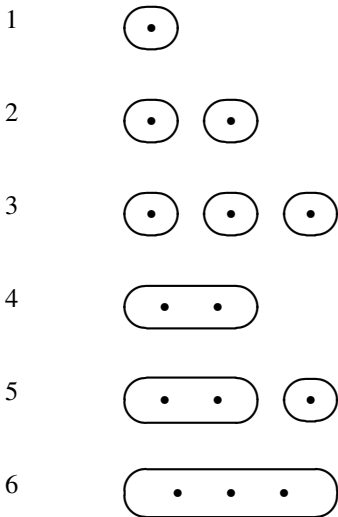
7. Effect of population structure

- ▶ D groups of N parents producing equal proportions of offspring (D assumed large)
- ▶ m : proportion of offspring that disperse uniformly before selection (Wright island model)
- ▶ $X_i(t)$: frequency of A in group i after dispersal in generation t
- ▶ $\overline{X(t)} = D^{-1} \sum_{i=1}^D X_i(t)$: frequency of A in the population in generation t
- ▶ $\overline{X(0)} = (ND)^{-1}$
- ▶ selection: within groups after dispersal

$$\tilde{x} = \frac{\sum_{t \geq 0} E_0 \left[\overline{X(t)^2(1 - X(t))} \right]}{\sum_{t \geq 0} E_0 \left[\overline{X(t)(1 - X(t))} \right]}$$

$$= \frac{\sum_{t \geq 0} P_0(A, A, B \text{ in the same group in generation } t)}{\sum_{t \geq 0} P_0(A, B \text{ in the same group in generation } t)}$$

States for the ancestors of 3 offspring



Transition matrix backward in time

Applying Möhle (1998) lemma

$$P^{\lfloor NDf_{22}^{-1}\tau \rfloor} \rightarrow \begin{pmatrix} I & 0 \\ F & 0 \end{pmatrix} \begin{pmatrix} e^{\tau A} & 0 \\ Be^{\tau A} & 0 \end{pmatrix}$$

as $D \rightarrow \infty$, where

$$F = \begin{pmatrix} f_{21} & f_{22} & 0 \\ 0 & f_{21} & f_{22} \\ f_{31} & f_{32} & f_{33} \end{pmatrix}$$

with f_{nk} the probability for n offspring in the same group to have k ancestors in different groups in the case of an infinite number of groups, and

$$A = \begin{pmatrix} 0 & 0 & 0 \\ 1 & -1 & 0 \\ 0 & 3 & -3 \end{pmatrix}$$

the infinitesimal generator for the **Kingman coalescent**.

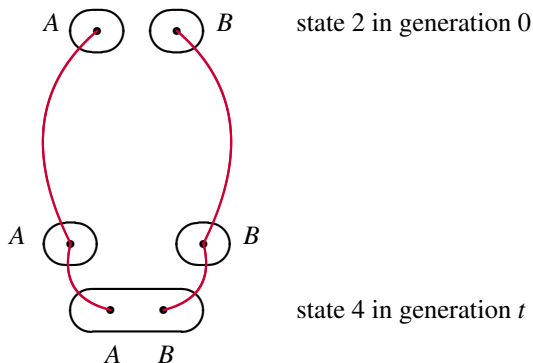
Two-time-scale argument for large D

- ▶ Fast scattering of lineages in the same group; slow collecting of lineages in different groups.
- ▶ Times spent with lineages in the same group can be neglected compared to times spent with lineages in different groups.
- ▶ The expected time spent with 2 lineages in different groups in number of generations is approximately NDf_{22}^{-1} , the effective population size.
- ▶ Moreover, the probability for two ancestors chosen at random in different groups in generation 0 to be of type A can be neglected.

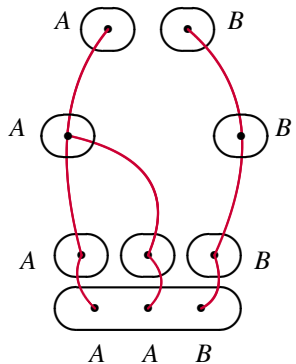
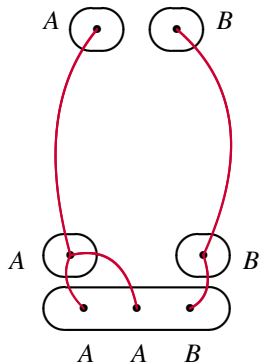
$$\sum_{t \geq 0} P_0(A, B \text{ in the same group in generation } t)$$

$$\approx \sum_{t \geq 0} P_{42}(t) P_0(A, B \text{ in different groups in generation } 0)$$

$$\approx f_{22} \times N D f_{22}^{-1} \times \overline{X(0)}(1 - \overline{X(0)})$$



$$\begin{aligned}
& \sum_{t \geq 0} P_0(A, A, B \text{ in the same group in generation } t) \\
& \approx \frac{1}{3} \sum_{t \geq 0} P_{62}(t) P_0(A, B \text{ in different groups in generation } 0) \\
& \approx \frac{1}{3} \times (f_{32} + f_{33}) \times NDf_{22}^{-1} \times \overline{X(0)}(1 - \overline{X(0)})
\end{aligned}$$



state 2 in generation 0

state 6 in generation t

$$\tilde{x} \approx \frac{1-f_{31}}{3(1-f_{21})} > \frac{1}{3}$$

with

$$f_{21} = \frac{(1-m)^2}{Nm(2-m) + (1-m)^2}$$

$$f_{31} = f_{21} \left[\frac{N(1-m) + 2(N-1)(1-m)^3}{N^2m(3-3m+m^2) + (3N-2)(1-m)^3} \right]$$

which means a **less stringent condition** for cooperation to be favored in the case of a **group-structured population**.

8. Variants and extensions

- Dispersal after selection:

$$\begin{aligned}u'(0) &= (b-d) \sum_{t \geq 0} E_0 \left[\overline{X(t)(1-X(t))} \right] \\ &+ (a-b-c+d) \sum_{t \geq 0} E_0 \left[\overline{X(t)^2(1-X(t))} \right] \\ &+ m(2-m)(b+c-2d) \sum_{t \geq 0} E_0 \left[\overline{X(t)^2} - \overline{X(t)}^2 \right] \\ &+ m(2-m)(a-b-c+d) \sum_{t \geq 0} E_0 \left[\overline{X(t)^3} - \overline{X(t)} \overline{X(t)^2} \right]\end{aligned}$$

$$\begin{aligned}
 E_0 \left[\overline{X(t)^2} - \overline{X(t)}^2 \right] &\approx P_0(A, B \text{ in different groups}) \\
 &- P_0(A, B \text{ in the same group}) \\
 E_0 \left[\overline{X(t)^3} - \overline{X(t)} \overline{X(t)^2} \right] &\approx P_0(A, A, B \text{ with } B \text{ in a different group}) \\
 &- P_0(A, A, B \text{ in the same group})
 \end{aligned}$$

$$\tilde{x} \approx \frac{(1-f_{31})(1-m)^2}{3(1-f_{21})} + \frac{2m(2-m)}{3} + \frac{(a-d)(N-1)^{-1}}{(a-b-c+d)} > \frac{1-f_{31}}{3(1-f_{21})}$$

which means an **even less stringent condition** for cooperation to be favored.

- Dispersal before selection with **exchangeable contributions of groups**:
 $D^{-\beta}$ the probability that a group at random provides a fraction ϕ of the migrants and every other a fraction $(1 - \phi)(D - 1)^{-1}$ in the case $\beta < 1$

$$u'(0) \approx \left[(b-d)f_{22} + \left(\frac{a-b-c+d}{3} \right) \left(f_{32} + f_{33} \frac{\mu_{32}}{\mu_{32} + \mu_{33}} \right) \right] \\ \times \mu_{21}^{-1} ND^{\beta} \times \overline{X(0)}(1 - \overline{X(0)})$$

where μ_{lk} is the **rate of transition from l to k lineages in different groups** backwards in time with ND^{β} generations as unit of time as $D \rightarrow \infty$

$$\tilde{x} \approx \frac{1 - f_{31} - f_{33} \frac{\mu_{33}}{\mu_{32} + \mu_{33}}}{3(1 - f_{21})} < \frac{1 - f_{31}}{3(1 - f_{21})}$$

which means a **more stringent condition** for cooperation to be favored.

Rates of transition for lineages in different groups

$$\mu_{lk} = \sum_{l \geq j \geq n \geq k} \lambda_{lj} p_{jn} f_{nk}$$

where

$$\lambda_{lj} = \binom{l}{j} (\phi m)^j (1 - \phi m)^l$$

$$p_{jn} = \frac{N^{[n]} S_j^{(n)}}{N^j}$$

with

$$N^{[n]} = N(N-1)\dots(N-n+1)$$

$$S_j^{(n)} = \frac{1}{n!} \sum_{\substack{j_1, \dots, j_n \geq 1 \\ j_1 + \dots + j_n = j}} \binom{j}{j_1, \dots, j_n}$$

f_{nk} is the probability of k types among n genes in an exact Wright-Fisher model with probability of mutation to a novel type m per gene per generation

$$f_{nk} = \sum_{j=0}^k \sum_{l=k-j}^{(n-1) \wedge (n-j)} \binom{n}{j} \frac{(Nm)^j (1-m)^{n-j} N^{[l]} S_{n-j}^{(l)}}{N^n - (1-m)^n N^{[n]}} f_{l,k-j}$$

where $|S_n^k| =$ coefficient of x^k in $x(x+1)\dots(x+n-1)$, in agreement with the approximation from the Ewens sampling formula for m small and N large

$$f_{nk} \approx \frac{M^k |S_n^k|}{M(M+1)\dots(M+n-1)}$$

where $M = 2Nm$.

Remark. The recursion equation can be extended to the Cannings model.

9. Summary and comments

- ▶ IPD in an **infinite population** can explain the spread of cooperation but only from a frequency $x > x^*$.
- ▶ IPD in a **finite population** can explain that cooperation is favored to go to fixation from a low frequency under the condition $x^* < \tilde{x}$.
- ▶ In a **large population**, $\tilde{x} \leq 1/3$ with a strict inequality leading to a more stringent condition if the contribution of parents in offspring is highly skewed.
- ▶ In a **group-structured population** with a large number of groups, the condition is less stringent with $\tilde{x} > 1/3$ unless the contribution of groups in offspring is skewed enough.

- ▶ The first-order effect of selection on the probability of fixation of an allele is given by the **projected average allelic effect on the selection coefficient**.
- ▶ This effect can be obtained under **more complex population assumptions**: interactions between any number of individuals, diploid populations with sex differences and overlapping generations, different timings for selection, mating and dispersal or local extinction in group-structured populations, etc.
- ▶ The **effect of selection of any order on the probability of fixation in a multiallele multilocus setting** can be obtained in a similar way in terms of expected sojourn times.

- ▶ The approximations given by the first-order effect of selection are valid **under very weak selection**, actually as long as the intensity of selection is small compared to the intensity of the other evolutionary forces, but **without constraints on the reproduction scheme**.
- ▶ An alternative approach under the assumption that the intensity of selection is of the same order of magnitude as the other evolutionary forces is a **diffusion approximation**, but then **the contributions of parents and groups in offspring cannot be too highly skewed**.

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Diffusion approximation for the Wright island model

- ▶ $s = \sigma(ND)^{-1}$ and dispersal before selection
- ▶ ND generations as unit of time as $D \rightarrow \infty$

Applying Ethier and Nagylaki (1980) conditions, the infinitesimal mean and variance of the limiting diffusion are

$$\begin{aligned}m(x) &= \sigma(a - b - c + d)x(1-x)(xf_{33} - x^*f_{22} + f_{21} - f_{31}) \\v(x) &= x(1-x)f_{22}\end{aligned}$$

and the probability of fixation of A from a small initial frequency x_0 is

$$\int_0^{x_0} \exp\left\{-2 \int_0^y \frac{m(x)}{v(x)} dx\right\} dy \approx x_0 + \sigma(a - b - c + d)x_0(1-x_0)f_{22} \left(\frac{1-f_{31}}{3f_{22}} - x^*\right)$$

Recursion equation for the Cannings model

$$f_{n,k} = \sum_{j=0}^k \sum_{l=k-j}^{(n-1) \wedge (n-j)} \sum_{\mathbf{r}} \frac{n! N^{[l]} m^j (1-m)^{n-j} E\left(\prod_{i=1}^l v_i^{[r_i]}\right)}{j! l! \left(\prod_{i=1}^l r_i!\right) N^{[n-j]} [1 - (1-m)^n E(\prod_{i=1}^n v_i)]} f_{l,k-j}$$

where $\mathbf{r} = (r_1, \dots, r_l)$ satisfy $r_1, \dots, r_l \geq 1$ with $r_1 + \dots + r_l = n - j$.