# Law of Large Numbers

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## The Model

- Suppose there is no latency period ( $T_1 = 0$  a.s.) and that  $\Delta T \sim \operatorname{Exp}(\alpha)$ . Consider a SIR model with constant population size equal to N. Let S(t) denote the number of susceptibles at time t, I(t) the number of infectious, R(t) the number of "removed" (i.e. "healed and immune").
- Hence the following equations, with  $P_1(t)$  and  $P_2(t)$  two standard mutually independent Poisson processes :

$$S(t) = S(0) - P_1 \left( \frac{\beta}{N} \int_0^t S(s) I(s) ds \right),$$

$$I(t) = I(0) + P_1 \left( \frac{\beta}{N} \int_0^t S(s) I(s) ds \right) - P_2 \left( \alpha \int_0^t I(s) ds \right)$$

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## The proportions

• Let us define  $s_N(t) = S(t)/N$ ,  $i_N(t) = I(t)/N$ . The equations for the proportions of susceptibles and infectious are written

$$s_N(t) = s_N(0) - \frac{1}{N} P_1 \left( \beta N \int_0^t s_N(r) i_N(r) dr \right),$$
  
$$i_N(t) = i_N(0) + \frac{1}{N} P_1 \left( \beta N \int_0^t s_N(r) i_N(r) dr \right) - \frac{1}{N} P_2 \left( \alpha N \int_0^t i_N(r) dr \right)$$

• Define the two martingales  $M_1(t) = P_1(t) - t$ ,  $M_2(t) = P_2(t) - t$ . We have

$$\begin{split} s_N(t) &= s_N(0) - \beta \int_0^t s_N(r) i_N(r) dr - \frac{1}{N} M_1 \left( \beta N \int_0^t s_N(r) i_N(r) dr \right), \\ i_N(t) &= i_N(0) + \beta \int_0^t s_N(r) i_N(r) dr - \alpha \int_0^t i_N(r) dr \\ &+ \frac{1}{N} M_1 \left( \beta N \int_0^t s_N(r) i_N(r) dr \right) - \frac{1}{N} M_2 \left( \alpha N \int_0^t i_N(r) dr \right). \end{split}$$

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Consider the process

$$\mathcal{M}_N(t) := rac{1}{N} M_1 \left( eta N \int_0^t s_N(r) i_N(r) dr 
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Let 
$$\mathcal{F}_t = \sigma\{s_N(r), i_N(r), 0 \le r \le t\}.$$

We have

#### Lemma

 $\{\mathcal{M}_N(t),\ t\geq 0\}$  is a  $\mathcal{F}_t$ -martingale which satisfies

$$\mathbb{E}\left[\mathcal{M}_{N}(t)\right] = 0, \quad \mathbb{E}\left[\left|\mathcal{M}_{N}(t)\right|^{2}\right] = \frac{\beta}{N}\mathbb{E}\int_{0}^{t} s_{N}(r)i_{N}(r)dr$$

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## Proof of the Lemma

- ullet For the martingale property, see the Notes. From this,  $\mathbb{E}[\mathcal{M}_N(t)]=0$ .
- Let  $0 = t_0 < t_1 < \cdots < t_n = t$ .

$$\begin{split} |\mathcal{M}_{N}(t)|^{2} &= \sum_{i=1}^{n} |\mathcal{M}_{N}(t_{i}) - \mathcal{M}_{N}(t_{i-1})|^{2} \\ &+ 2 \sum_{1 \leq i < j \leq n} [\mathcal{M}_{N}(t_{i}) - \mathcal{M}_{N}(t_{i-1})] [\mathcal{M}_{N}(t_{j}) - \mathcal{M}_{N}(t_{j-1})] \end{split}$$

$$\mathbb{E}\left(|\mathcal{M}_N(t)|^2\right)^2 = \sum_{i=1}^n \mathbb{E}\left(|\mathcal{M}_N(t_i) - \mathcal{M}_N(t_{i-1})|^2\right)$$

• When  $|t_i - t_{i-1}| \to 0$ ,

$$\sum_{i=1}^{n} |\mathcal{M}_{N}(t_{i}) - \mathcal{M}_{N}(t_{i-1})|^{2} \to \sum_{0 \le r \le t} |\Delta \mathcal{M}_{N}(r)|^{2}$$
$$= N^{-2} P_{1} \left(\beta N \int_{0}^{t} s_{N}(r) i_{N}(r) dr\right)$$

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$$\mathbb{E}\left[\left|\mathcal{M}_{N}(t)\right|^{2}\right] = N^{-2}\mathbb{E}P_{1}\left(\beta N \int_{0}^{t} s_{N}(r)i_{N}(r)dr\right)$$
$$= \frac{\beta}{N}\mathbb{E}\int_{0}^{t} s_{N}(r)i_{N}(r)dr.$$

- This implies that  $\sup_{0 \le t \le T} |\mathcal{M}_N(t)| \to 0$  in probability. In fcat this is true a.s.
- This follows from

$$\sup_{0 \le t \le T} \left| \frac{P_1(Nt)}{N} - t \right| \to 0 \text{ a.s. as } N \to \infty$$

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• We prove the above results in three steps. Let P(t) be a rate  $\lambda$  Poisson process

$$n^{-1}P(n) = n^{-1}\sum_{i=1}^{n}[P(i) - P(i-1)]$$
  
 $\rightarrow \lambda \text{ a.s. as } n \rightarrow \infty.$ 

Second step

$$t^{-1}P(t) = t^{-1}P([t]) + t^{-1}(P(t) - P([t]))$$
$$|t^{-1}P(t) - \lambda| \le |t^{-1}P([t]) - \lambda| + t^{-1}P([t] + 1) - t^{-1}P([t]).$$

• We have just proved that  $N^{-1}P(Nt) \to \lambda t$  a.s. for all t > 0. We have a sequence of increasing functions which converges to a continuous function, hence by the second Dini theorem, the convergence is uniform.

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• We have the Law of Large Numbers :

$$\sup_{0 \leq t \leq \mathcal{T}} \{|s_{\mathcal{N}}(t) - s(t)| + |i_{\mathcal{N}}(t) - i(t)|\} 
ightarrow 0$$
 a.s., where

(s(t), i(t)) solves

$$\begin{cases} \frac{ds}{dt}(t) = -\beta s(t)i(t), \ t > 0, \\ \frac{di}{dt}(t) = \beta s(t)i(t) - \alpha i(t), \ t > 0. \end{cases}$$

• Define 
$$X(t) = \begin{pmatrix} s(t) \\ i(t) \end{pmatrix}$$
,  $X_N(t) = \begin{pmatrix} s_N(t) \\ i_N(t) \end{pmatrix}$ ,  $\overline{X}_N(t) = X(t) - X_N(t)$ ,  $Y_N(t) = \begin{pmatrix} \mathcal{M}_N(t) \\ \mathcal{N}_N(t) - \mathcal{M}_N(t) \end{pmatrix}$ , and finally  $F\begin{pmatrix} x \\ y \end{pmatrix} = \begin{pmatrix} -\beta xy \\ \beta xy - \alpha y \end{pmatrix}$ . For  $0 \le x, y, x', y' \le 1$ ,

$$\left\| F \begin{pmatrix} x \\ y \end{pmatrix} - F \begin{pmatrix} x' \\ y' \end{pmatrix} \right\| \le C(\alpha, \beta) \left\| \begin{pmatrix} x \\ y \end{pmatrix} - \begin{pmatrix} x' \\ y' \end{pmatrix} \right\|$$

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• Define  $X(t) = {s(t) \choose i(t)}$ ,  $X_N(t) = {s_N(t) \choose i_N(t)}$ ,  $\overline{X}_N(t) = X(t) - X_N(t)$ ,  $Y_N(t) = {M_N(t) \choose \mathcal{N}_N(t) - \mathcal{M}_N(t)}$ , and finally  $F {x \choose y} = {-\beta xy \choose \beta xy - \alpha y}$ . For  $0 \le x, y, x', y' \le 1$ ,

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$$\overline{X}_N(t) = \overline{X}_N(0) + \int_0^t [F(X(r)) - F(X_N(r))] dr + Y_N(t).$$

- We have proved that  $\sup_{0 \le t \le T} \|Y_N(t)\| \to 0$  a.s. as  $N \to \infty$
- Let  $\varepsilon_N(t) = \sup_{0 \le r \le t} \|Y_N(r)\|$ . We have

$$\|\overline{X}_N(t)\| \le \|\overline{X}_N(0)\| + C(\alpha, \beta) \int_0^t \|\overline{X}_N(r)\| dr + \varepsilon_N(t)$$

It then follows from Gronwall's Lemma that

$$\sup_{0 \le r \le t} \|\overline{X}_N(r)\| \le (\|\overline{X}_N(0)\| + \varepsilon_N(t)) \exp(C(\alpha, \beta)t),$$

hence the result, provided that  $.\overline{X}_N(0) \to 0$ 

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