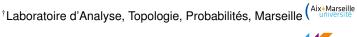
# **Unscented Variational Bayesian Inference**

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Journée GdR ISIS "Traitement du signal de données à valeurs complexes", 12 juin 2013

## CONTEXT

#### The phase Retrieval Problem

```
Given b \in \mathbb{C}^m, find x \in \mathbb{C}^n such that |Ax| = b,
```

where A is a matrix in  $\mathbb{C}^{m \times n}$ .

#### Audio Applications

- Source separation [Sturmel, 2011]
- Sound transformation in the time-frequency domain [Olivero, 2012]
- Synthesis via scattering coefficients [Mallat, 2012]

#### Solver for the phase retrieval problem

- Iterative algorithms like [Griffin and Lim, 1984]
- Recent algorithms used convex relaxation techniques and semi-definite programming [Candes, 2011]

#### Our contribution

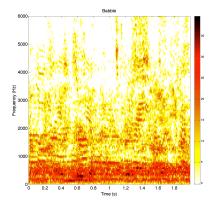
- Bayesian model
- Unscented transform

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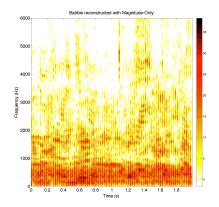
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#### EXAMPLE IN THE GABOR CASE WITH A BABBLE SIGNAL

#### Signal reconstructed via Ax



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## PHASE RETRIEVAL PROBLEM

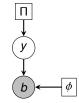
Given 
$$b \in \mathbb{R}^m_+$$
, find  $x \in \mathbb{C}^n$  such that  $|Ax| = b$ ,

where *A* is a matrix in  $\mathbb{C}^{m \times n}$ .

Find 
$$Ax = y \in \mathbb{C}^m$$
 such that 
$$\begin{cases} |y| = b \\ (I - AA^{\dagger}) \\ y = 0 \end{cases}$$

For noisy y and b, the model reads :

$$p(y) \propto exp(-||\Pi y||^2), \quad y \in \mathbb{C}^m$$
$$p(b_m | y_m, \phi_m) \propto exp\left\{-\phi_m \left(|y_m| - b_m\right)^2\right\}$$
If  $\phi_m = \phi \in \mathbb{R}_+$ , this model amounts to
$$\min_y ||\Pi y||^2 + \phi |||y| - b||^2$$



## PHASE RETRIEVAL PROBLEM

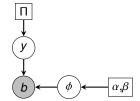
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$$p(\phi_m) \propto \text{Gamma}(\alpha, \beta)$$



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## CONTENT

Variational Bayesian Inference

**Unscented Variational Bayesian Inference** 

Experiments Logistic Regression

Conclusion

- ► *H* : hidden random variables
- ► V : visible random variables (i.e. observations)
- M : model defining the probabilistic dependencies between H and V

Goal : evaluate the model posterior probablity

$$p(\mathcal{M}|V = v) \propto p(V = v|\mathcal{M})\pi(\mathcal{M})$$

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which comes down to evaluate the log-evidence

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Now, we will drop the subscript  $\ensuremath{\mathcal{M}}$ 

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## VARIATIONAL BAYESIAN INFERENCE

$$\mathcal{E}(v) = \ln p(V = v)$$
  
=  $\ln \sum_{h} p(V = v, H = h)$   
=  $\ln \sum_{h} q(h) \frac{p(V = v, H = h)}{q(h)}$ , for any  $q \in Q$   
$$\geq \underbrace{\mathbb{E}_{H \sim q} \{\ln P(V = v, H)\}}_{\mathcal{F}_{v}(q)} + \underbrace{\mathbb{E}_{H \sim q} \{-\ln q(H)\}}_{\text{Entropy}: \mathcal{G}(q)}$$

 $\mathcal{L}_{v}(q)$  is a lower bound of the evidence (ELBO).

$$\mathcal{E}(v) = \max_{q \in Q} \mathcal{L}_v(q),$$

The optimal distribution is equal to the posterior distribution p(H|V), but ....

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## VARIATIONAL BAYESIAN INFERENCE

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- 1. *Q* is the family of all probabilities of distributions on *H*.
- 2. How can we compute  $\mathcal{F}_{v}(q)$  and  $\mathcal{G}(q)$ ?
  - The popular Mean-Field approximation restricts Q to factorized distributions [Jordan and al, 1999].
  - [Gershman, 2012] proposed a new method is also called "Nonparametric Variational Inference".

## NONPARAMETRIC VARIATIONAL INFERENCE

$$\mathcal{E}(v) = \max_{q \in \mathcal{Q}} \mathcal{L}_{v}(q) = \underbrace{\mathbb{E}_{H \sim q} \{ \ln P(V = v, H) \}}_{\mathcal{F}_{v}(q)} + \underbrace{\mathbb{E}_{H \sim q} \{ -\ln q(H) \}}_{\text{Entropy}: \mathcal{G}(q)}$$

▶ [Gershman, 2012]

1. 
$$q(h) = \frac{1}{N} \sum_{n=1}^{N} \frac{1}{(2\pi\sigma_n^2)^{D/2}} \exp\left\{-\frac{1}{2\sigma_n^2} \|h - \mu_n\|^2\right\}.$$

- 2. In  $P(V = v, \cdot)$  is obtained with a second order approximation.
- 3. The entropy is lowly bounded.
- We build upon and extend this approach :
  - 1. we will consider a non isotropic mixture of gaussians
  - 2. we use unscented transform to evaluate the expectations.

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## UNSCENTED TRANSFORM

#### [Julier and Uhlmann, 1995]

Let *X* be a random variable with  $\mathbb{E}[X] = \overline{X}$  and  $cov(X) = RR^T$  and *f* a twice differentiable function :

$$Z = f(X) \approx f(\overline{X}) + \nabla_{\overline{X}} f^{T} (X - \overline{X}) + \frac{1}{2} (X - \overline{X})^{T} H f_{\overline{X}} (X - \overline{X})$$
$$\mathbb{E}[f(X)] \approx f(\overline{X}) + 0 + \frac{1}{2} Tr \{\mathbb{E}[(X - \overline{X})^{T} (X - \overline{X})] H f_{\overline{X}}\}$$
(1)
$$= f(\overline{X}) + \frac{1}{2} Tr \{cov(X) H f_{\overline{X}}\}$$

For all random variable with mean  $\overline{X}$  and covariance cov(X), we obtain a second-order approximation of  $\mathbb{E}[f(X)]$ .

 We will construct a discrete random variable χ with values χ<sub>i</sub> and probabilities ω<sub>i</sub>.

$$(1) a^T b = Tr(a^T b) = Tr(ba^T)$$

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#### UNSCENTED TRANSFORM

Let X be a random variable with  $\mathbb{E}[X] = \overline{X}$  and  $cov(X) = RR^{T}$ .

A set of sigma points  $\{\chi_i\}$  and weights  $\{\omega_i\}$  has the same mean and covariance.

$$\chi_{0} = \overline{X} \qquad \qquad \omega_{0} = \frac{\kappa}{n+\kappa}$$

$$\chi_{i} = \overline{X} + \sqrt{n+\kappa}R_{i} \qquad \qquad \omega_{i} = \frac{1}{2(n+\kappa)}$$

$$\chi_{i+n} = \overline{X} - \sqrt{n+\kappa}R_{i} \qquad \qquad \omega_{i+n} = \frac{1}{2(n+\kappa)}, \quad i = 1, .., n$$

$$R_{i} \text{ is the i-th column of } R$$

#### Proposition

If *f* is twice differentiable, then  $\mathbb{E}[f(\chi)] = \sum_{i=0}^{2n} \omega_i f(\chi_i)$  is a second order approximation of  $\mathbb{E}[f(X)]$ .

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## UNSCENTED TRANSFORM

Let *X* be a random variable with  $\mathbb{E}[X] = \overline{X}$  and  $cov(X) = UU^T + VV^T$ . We define a set of *sigma points* :  $\{X_i^u, X_j^v\}$  with corresponding weights  $\{\omega_i^u, \omega_j^v\}$ .

For  $i = 1, ..., n_u, j = 1, ..., n_v$ :

$$X_{i}^{u} = \begin{cases} \overline{X} \\ \overline{X} + \alpha_{i}^{u} u_{i} \\ \overline{X} - \alpha_{i+n}^{u} u_{i} \end{cases} \text{ and } X_{j}^{v} = \begin{cases} \overline{X} \\ \overline{X} + \alpha_{j}^{v} v_{j} \\ \overline{X} - \alpha_{j+n}^{v} v_{j} \end{cases}$$
$$\text{with } \sum_{i} \omega_{i}^{u} + \sum_{j} \omega_{j}^{v} = 1$$

#### Proposition

If *f* is twice differentiable, then  $\mathbb{E}[f(\chi)] = \sum_{k} \omega_k f(\chi_k)$  is a second order approximation of  $\mathbb{E}[f(X)]$ .

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## UNSCENTED VARIATIONAL BAYESIAN INFERENCE

$$\mathcal{E}(v) = \max_{q \in Q} \mathcal{L}_{v}(q) = \underbrace{\mathbb{E}_{H \sim q} \{ \ln P(V = v, H) \}}_{\mathcal{F}_{v}(q)} + \underbrace{\mathbb{E}_{H \sim q} \{ -\ln q(H) \}}_{\text{Entropy} : \mathcal{G}(q)}$$

In our approach,

1. The family of distributions is described with

$$q(h) = \frac{1}{N} \sum_{n=1}^{N} \frac{1}{\sqrt{(2\pi)^{D} |\Sigma_{n}|}} \exp\left\{-\frac{1}{2}(x-\mu)^{\mathsf{T}} \Sigma_{n}^{-1}(x-\mu)\right\}$$
  
where 
$$\begin{cases} \Sigma_{n} = \sigma_{n}^{2} I + W_{n} W_{n}^{\mathsf{T}} \\ W_{n} \in \mathbb{R}^{p \times D} \text{ with } p << D \end{cases}$$

2. The **unscented transform** gives a second order approximation of the expectations.

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#### LOGISTIC REGRESSION : EXPERIMENTS

- Obervations :  $V = \{x_t, c_t\}_{t=1}^T$ , where  $x_t \in \mathbb{R}^D$ ,  $c_t \in \{\pm 1\}$ .
- ▶ Hidden variables :  $H = \{w, \alpha\}$ , where  $w \in \mathbb{R}^D$ ,  $\alpha \in \mathbb{R}_+$

The probabilities describing the possible outcomes are modeled using a logistic function :

$$p(c_t|w,x_t) = \frac{1}{1+e^{-c_tw^Tx_t}}$$

The regression parameters are given with probabilities :

$$p(w|\alpha) = g(w; 0, \alpha^{-1})$$
$$p(\alpha) = \gamma(\alpha; a, b)$$

We compare our approach with the one proposed in [Gershmann, 2012] for simulated data and UCI datasets

Comparable results

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## CONCLUSION

We show a promising new framework for Variational Bayesian Inference, with few hypothesis on the joint probability involved (twice differentiability).

Advantages :

- Flexibility
- Confidence intervals

Drawbacks :

Computational cost

Future directions :

- Estimation of non-uniform mixture coefficients
- ► Situations where the function is L-Lipschitz differentiable
- Model selection
- Intensive experiments for the phase retrieval problem

#### AKNOWLEDGEMENTS

- ANR Project ASAP
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Thanks for your attention